

# Genomic Analyses of Visual Cognition: Perceptual Rivalry and Top-Down Control

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Visual cognition in humans has traditionally been studied with cognitive behavioral methods and brain imaging, but much less with genetic methods. Perceptual rivalry, an important phenomenon in visual cognition, is the spontaneous perceptual alternation that occurs between two distinct interpretations of a physically constant visual stimulus (e.g., binocular rivalry stimuli) or a perceptually ambiguous stimulus (e.g., the Necker cube). The switching rate varies dramatically across individuals and can be voluntarily modulated by observers. Here, we adopted a genomic approach to systematically investigate the genetics underlying binocular rivalry, Necker cube rivalry and voluntary modulation of Necker cube rivalry in young Chinese adults (*Homo sapiens*, 81% female, 20 ± 1 years old) at multiple levels, including common single nucleotide polymorphism (SNP)-based heritability estimation, SNP-based genome-wide association study (GWAS), gene-based analysis, and pathway analysis. We performed a pilot GWAS in 2441 individuals and replicated it in an independent cohort of 943 individuals. Common SNP-based heritability was estimated to be 25% for spontaneous perceptual rivalry. SNPs rs184765639 and rs75595941 were associated with voluntary modulation, and imaging data suggested genotypic difference of rs184765639 in the surface area of the left caudal-middle frontal cortex. Additionally, converging evidence from multilevel analyses associated genes such as *PRMT1* with perceptual switching rate, and *MIR1178* with voluntary modulation strength. In summary, this study discovered specific genetic contributions to perceptual rivalry and its voluntary modulation in human beings. These findings may promote our understanding of psychiatric disorders, as perceptual rivalry is a potential psychiatric biomarker.

**Key words:** GWAS; heritability; perceptual rivalry; structural MRI; voluntary modulation

## Significance Statement

Perceptual rivalry is an important visual phenomenon in which our perception of a physically constant visual input spontaneously switches between two different states. There are individual variations in perceptual switching rate and voluntary modulation strength. Our genomic analyses reveal several loci associated with these two kinds of variation. Because perceptual rivalry is thought to be relevant to and potentially an endophenotype for psychiatric disorders, these results may help understand not only visual cognition, but also psychiatric disorders.

## Introduction

Visual cognition is traditionally studied in animals with electrophysiology and in humans with behavioral and brain imaging methods. Genomic technologies including microarray and se-

quencing provide a powerful approach to study human visual cognition. Here, we chose to investigate perceptual rivalry because it is an intriguing visual cognitive phenomenon known for centuries as well as a potential biomarker for psychiatric disor-

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neous perceptual alternation may be mediated by brain regions in visual cortex (Meng et al., 2005; Frässle et al., 2014; Ishizu and Zeki, 2014) and frontoparietal networks (Cohen, 1959; Ricci and Blundo, 1990; Meenan and Miller, 1994; Bisiach et al., 1999; Miller et al., 2000; Slotnick and Yantis, 2005; Ge et al., 2007; Fagard et al., 2008; Shen et al., 2009; Sterzer et al., 2009; Kanai et al., 2010; Knapen et al., 2011; Wang et al., 2013; Weilhhammer et al., 2013; Megumi et al., 2015

ders (Ngo et al., 2011). When observing a perceptual rivalry stimulus, our brain oscillates spontaneously between two distinct percepts although the retinal input remains unchanged. Well-studied examples of perceptual rivalry include the Necker cube (Fig. 1A, middle; Necker, 1832), which can be perceived as either facing top-left (Fig. 1A, left) or bottom-right (Fig. 1A, right), and binocular rivalry, which can be elicited when two eyes receive different sensory inputs (Fig. 1B). Perceptual rivalry can be top-down influenced: individuals may change their switching rate via voluntary effort (Washburn and Gillette, 1933; Pelton and Solley, 1968; Liebert and Burk, 1985; Strüder and Stadler, 1999; Toppino, 2003; Meng and Tong, 2004; van Ee et al., 2005; Kornmeier et al., 2009; Scocchia et al., 2014). Moreover, individuals with clinical conditions such as bipolar disorder (Miller et al., 2003; Nagamine et al., 2009; Vierck et al., 2013) and autism (Ropar et al., 2003; Wimmer and Doherty, 2010; Allen and Chambers, 2011; Robertson et al., 2013) display different rivalry patterns; slower switching rate in bipolar disorder and longer mixed perceptual states in autism.

The physiological bases for perceptual rivalry and top-down voluntary modulation have been investigated with electrophysiology and imaging methods in monkeys (Leopold and Logothetis, 1996; Xu et al., 2016) and with electroencephalography (Strüder et al., 2001; Kornmeier and Bach, 2004; Isoğlu-Alkac and Strüder, 2006; Mathes et al., 2006; Pitts et al., 2008), magnetoencephalography (Parkkonen et al., 2008), magnetic resonance imaging (MRI; Lumer et al., 1998; Inui et al., 2000; Sterzer and Kleinschmidt, 2007; Watanabe et al., 2014), and brain stimulation (Ngo et al., 2013; Brascamp et al., 2018) in humans. Sponta-

## Materials and Methods

### Participants

The GWAS discovery and replication cohorts consisted of college students recruited from the Chongqing Medical University, with a mean age of 20 years ( $SD = 1$ ), 81% female, 94% Han. The psychiatric history and medication history were screened by medical examination at the time of college entrance and by self-reported questionnaires just before the current study. None of the subjects reported having psychiatric disorders, having a history of brain injuries, having a first-degree relative with psychiatric disorders, having vision disorders, or currently taking medication. Visual acuity for each subject was measured with the software program FrACT v3.8.2, which is based on the Freiburg Vision Test (Bach, 1996). A Landolt-C, as an opto-type, was presented at the center of the screen, 1.4 m away from subjects. Subjects were asked to make an eight-alternative forced-choice judgment on the orientation of the Landolt-Cs. The program estimated acuity threshold using a method called Best Parameter Estimation by Sequential Testing. All subjects had normal or corrected-to-normal visual acuity. No significant difference in ethnicity or place of origin between the discovery and replication cohorts was found (Pearson's  $\chi^2$  test). Data for the discovery and replication cohorts were collected in the autumn of 2014 and the spring of 2015, respectively. Written informed consents were collected before experiments. The human subject review committee in the School of Psychological and Cognitive Sciences at the Peking University approved the experimental procedures.

### Experimental design

In the binocular rivalry experiment, the stimulus was presented on a dark background and consisted of two overlapped sinusoidal gratings (orientation: horizontal and vertical; color: red and blue; diameter:  $1.5^\circ$  of visual angle; spatial frequency: 5 cycles/ $^\circ$ ; phase: random; contrast: 0.9). The two gratings drifted simultaneously in vertical or horizontal directions at a speed of  $1^\circ/s$  (Fig. 1*B*). A purple dot at the center of the gratings served as a fixation point. Subjects viewed the stimulus through red-blue anaglyph eyeglasses, with the red horizontal grating presented to their left eye and the blue vertical grating presented to their right eye. Since the two small gratings presented in the fovea were dramatically different in color, orientation, and moving direction, participants experienced clear-cut binocular rivalry with little mixed percept; their percept alternated between a blue leftward-moving grating and a red upward-moving grating. For the Necker cube experiment, a cube composed of 12 white bars was presented on a dark background, subtending  $1^\circ \times 1^\circ$  of visual angle. A red dot at the center of the cube served as a fixation point. To minimize eye movement (Ellis and Stark, 1978), subjects were requested to fixate the central dot throughout the experiment.

All stimuli were presented on a Lenovo LCD monitor with a spatial resolution of  $1024 \times 768$  and a refresh rate of 60 Hz. Luminance was measured with a MINOLTACS-100A Chroma Meter (Konica Minolta Sensing Americas). The luminance levels of the background, the gratings, and the Necker cube were  $0\text{--}1\text{ cd/m}^2$ ,  $46\text{ cd/m}^2$  (mean luminance), and  $200\text{ cd/m}^2$ , respectively. The stimuli were generated and controlled using MATLAB and Psychtoolbox (Brainard, 1997).

At the beginning, subjects were informed about the reversibility of the stimuli and two possible perceptual states (Fig. 1). After they were able to perceive both states, they took 1 min training. Subjects were instructed to press one key when they perceived the red horizontal grating (or the cube facing top-left) and the other key for the blue vertical grating (or the cube facing bottom-right). Subjects were notified about the possibility of mixed percept during rivalry and they were instructed to classify mixed percept into either state based on which state covered the majority of the stimulus field. Subjects were asked to report occurrence of persistent mixed percept after each trial. Eleven subjects who reported persistent mixed percept were excluded from the experiment. For the binocular rivalry experiment, subjects were instructed to view the stimulus naturally without any voluntary effort. The experiment consisted of eight trials, each lasting 1 min, with a 1–3 min (self-paced) break in between. Subjects decided when to start the next trial by pressing a “start” button that would appear on the screen 1 min after the end of a trial. For the Necker cube experiment, there were two conditions: the spontaneous

switch condition and the voluntary modulation condition. In the former condition, subjects viewed the stimulus passively without any voluntary effort. In the latter condition, subjects were asked to do their best to increase their perceptual switching rate. The spontaneous switch condition was followed by the voluntary modulation condition. Each condition consisted of five trials, each lasting 1 min, with a 1–3 min break in between. Perceptual switching rate was measured as the number of perceptual switches per minute, averaged across all the trials in the spontaneous switch condition (8/5 trials for binocular/Necker cube rivalry). The voluntary modulation strength was calculated as (the switching rate in the voluntary modulation condition – the switching rate in the spontaneous switch condition)/the switching rate in the spontaneous switch condition. We did not ask subjects to voluntarily increase their perceptual switching rate in the binocular rivalry experiment. This is because, in a pilot study, we found such voluntary modulation had little effect on binocular rivalry, which is consistent with previous findings (Meng and Tong, 2004). In the voluntary modulation condition for the Necker cube, subjects were asked to speed up, rather than slow down, their perceptual switching rate. We chose this option because we were concerned that instructing subjects to slow down switching rate might be confounded by factors such as simply diverting attention from the task (but see Discussion).

### Genotyping and quality control

**Discovery cohort.** DNA was extracted from peripheral blood of participants using the QuickGene whole-blood genome DNA extract system (Kurabo Industries), and was genotyped for 894,517 common SNPs using the HumanOmniZhongHua-8 Beadchip v1.2 (Illumina). Common quality control parameters were applied to the genotype data and retained 830,937 SNPs. SNPs were included in the analysis if they met the following criteria: call rate  $>0.95$ , minor allele frequency (MAF)  $>0.01$ , and Hardy–Weinberg equilibrium test with  $p > 10^{-4}$ . Differences in allele frequencies between the discovery and the replication cohorts were examined using paired sample  $t$  test, and no significant deviation was observed (MAF of the replication cohort were all within 2 SD from the mean MAF of the discovery cohort). Individuals with genotyping call rate  $<0.95$  were excluded. Potential duplicates or close relatives were examined by calculating an identity-by-state (IBS) similarity matrix implemented in PLINK (Purcell et al., 2007), and none was excluded due to IBS distance  $\leq 0.75$  (corresponding to siblings). Individuals with extreme phenotypes outside 4 SD from the population mean were taken as phenotypic outliers and removed from subsequent analyses (Table 1-1, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t1-1>). Population stratification within this cohort was examined with EIGENSTRAT (Price et al., 2006) via principal component analysis (PCA). Genetic outliers were detected and excluded automatically under a default mode [the maximum number of outlier removal iterations was 5; an individual outside 6 SD from the population mean in terms of genetic principal component (PC) was removed as an outlier]. Approximately 1.5% of participants in the discovery cohort were excluded as outliers (phenotypic and genetic). The sex ratios of the outliers were similar to those of the full sample. After outlier removal, the genetic backgrounds of the ethnic groups in the remaining populations did not significantly differ from that of the Han individuals. A total of 2441 (2097 for binocular rivalry, 1880 for Necker cube rivalry, 1876 for voluntary modulation strength) subjects with both phenotypic and genotypic data available were included in the final analysis (Table 1-1, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t1-1>).

**Replication cohort.** The replication cohort was a new sample genotyped on the HumanOmniZhongHua-8 Beadchip v1.1. The same quality control protocol as that at the discovery stage was applied, resulting in a final dataset of 830,937 SNPs genotyped in a homogenous sample of 943 individuals (870 for binocular rivalry, 748 for Necker cube rivalry, 744 for voluntary modulation strength), with 3.5% of the replication cohort identified as outliers and discarded (Table 1-1, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t1-1>).

### Statistical analysis

**Heritability estimation.** We performed a restricted maximum likelihood (REML) analysis by GCTA v1.24 (Yang et al., 2011) to estimate the phe-

notypic variance explained by all common SNPs (SNP heritability). This method measures the variance in the trait that is attributable to the genetic difference across the population. The discovery and replication cohorts were combined into one sample. We estimated the genetic relationship matrix (GRM) using all autosomal markers that were genotyped and checked for quality control as described above (altogether 830,937). One individual in each pair was excluded if the estimated genetic relatedness of that pair was  $>0.025$  (Yang et al., 2010). Twenty-four individuals were excluded and 2935 remained. We adopted the GRM-REML (GREML) method to estimate the variance explained by common SNPs, with the first 20 eigenvectors from PCA included as covariates. We used an additive model because dominant effects accounted for very little of the total genetic effects when SNP heritability analysis assumed a full model. We used the online GCTA-GREML Power Calculator (Visscher et al., 2014) to calculate the statistical power. We estimated genetic correlations for binocular rivalry and Necker cube rivalry by bivariate GREML implemented in GCTA.

We explored the genetic correlations between the measured phenotypes in our Chinese population and five psychiatric disorders (bipolar disorder, major depression, schizophrenia, autism spectrum disorder, and attention deficit hyperactivity disorder) in a Caucasian population by linkage disequilibrium (LD) score regression analysis (Bulik-Sullivan et al., 2015) via the web-based LD Hub (Zheng et al., 2017). Summary GWAS data from the discovery cohort were used as inputs. Significance threshold was set at  $p = 0.0033$ .

**Imputation.** Genotypes were pre-phased into haplotypes with SHAPEIT (Delaneau et al., 2012; Howie et al., 2012). Imputation was then performed using IMPUTE v2.3.1 (Howie et al., 2009) based on 1000 Genomes haplotype data [Phase I integrated variant set release (SHAPEIT2) in NCBI build 37/UCSC hg19] with 36,820,992 SNPs, 1384,273 short biallelic indels and 14,017 structural variations. Quality control retained those variants with information (INFO)  $>0.4$  and MAF  $>0.01$  for subsequent analyses, which were 7,992,300; 7,990,224; and 7,989,184 SNPs for binocular rivalry, Necker cube rivalry, and voluntary modulation strength, respectively.

**GWAS.** Perceptual switching rates were taken as quantitative traits. The first 10 PCs estimated by EIGENSTRAT were tested for phenotypic associations using linear regression. Effects of demographic factors (sex, age, ethnicity, and birthplace) were also tested using linear regression with phenotypes. None of those factors showed association with  $p < 0.05$  when all factors were included as independent variables or only a particular factor was included as an independent variable. Thus, no covariate was included in subsequent association tests. Based on the Score Method in the Frequentist Test framework, association tests of imputed SNPs were performed with SNPTEST v2.5 (Marchini et al., 2007), assuming an additive model. The full set of  $p$  values that emerged from association analyses was loaded and visualized in Haploview v4.2 (Barrett et al., 2005) to generate Manhattan plots. Basic statistical analyses were performed, the genomic inflation factor  $\lambda$  was calculated, and quantile–quantile (Q–Q) plots for observed versus expected  $p$  values were generated, all by R v3.2.1 (<https://www.R-project.org>). Genome-wide significance threshold was set at  $(5/3) \times 10^{-8}$ . Regional association plots were drawn for two mega-bases around top candidate SNPs from the discovery GWAS by the web-based LocusZoom (<http://locuszoom.sph.umich.edu/locuszoom>), only with variants that passed quality control. Quanto v1.2 was used for power calculation (Gauderman and Morrison, 2006). The power of this study to discover those genome-wide significant SNPs was  $>0.85$  for Necker cube rivalry and its voluntary



**Table 1. Summary of phenotypes in all cohorts**

	Sex, male%	Age, y	BR	NC	VOLUNTARY
Discovery cohort					
N		2441	2097	1880	1876
Range		16~25	0.63~72.38	1.30~21.60	−0.90~4.68
Mean ± SD	21.2	20 ± 1	29.73 ± 9.71	7.74 ± 3.04	0.54 ± 0.78
Replication cohort					
N		943	870	748	744
Range		18~26	0.88~61.88	0~63.50	−6.00~24.90
Mean ± SD	12.4	21 ± 1	29.61 ± 9.15	7.72 ± 2.79	0.53 ± 0.75

All data are untransformed raw phenotypic scores. BR, binocular rivalry; NC, Necker cube rivalry; VOLUNTARY, voluntary modulation strength; N, number of subjects. For more details on studied cohorts, see Table 1-1 (available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t1-1>).

### Heritability estimation

We adopted the GREML method to estimate SNP heritability of these perceptual rivalry traits, through quantifying the proportion of phenotypic variance explained by all genotyped autosomal SNPs. Results showed that ~25% (SE = 13%,  $p = 0.026$ ) of the phenotypic variance in BR could be explained by all genotyped common SNPs, whereas that percentage for NC was 23% (SE = 16%,  $p = 0.069$ ; Table 2). The genetic correlation between BR and NC was 0.24 (SE = 0.46). This SNP-based analysis suggests a moderate SNP heritability for spontaneous perceptual rivalry. BR and NC might share partial genetic basis. No significant common SNP heritability estimate was found for VOLUNTARY (Table 2). LD regression did not find any significant genetic correlation with the tested psychiatric disorders (Table 2-1, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t2-1>).

### GWAS of single-marker association

Classic GWAS was first performed in the discovery cohort on the imputed markers after quality control. The genomic inflation factor  $\lambda$  was 1.00 for all association tests. Thus, the significance reported herein was not affected by population stratification. The overall association results across the whole genome are shown in Manhattan plots (Fig. 2). One, 15, and 92 SNPs reached genome-wide significance after multiple-testing correction ( $p < 1.67 \times 10^{-8}$ ) for BR, NC, and VOLUNTARY, respectively (Fig. 2; Table 3-1, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t3-1>). The number of markers with genome-wide significance was much larger than that expected by chance, as illustrated by Q-Q plots (Fig. 3). The only genomic significant SNP for BR was rs11742135 (chr5:121646077A>C; Fig. 4A), but its association signal in the replication cohort was not significant ( $p = 0.6$ ). For NC, one genome-wide significant SNP, rs190906337 (chr2:211072859T>C), reached an uncorrected  $p = 0.018$  at the replication stage (Fig. 4B). Two SNPs for VOLUNTARY, namely, rs75595941 (chr6:68772523T>C; Fig. 4C) and rs184765639 (chr4:79651292G>A; Fig. 4D), displayed significant association signals after multiple-testing correction ( $p < 4.63 \times 10^{-4}$ ) at the replication stage (Table 3). The effect directions for these markers in the replication cohort were consistent with those in the discovery cohort. rs184765639 is predicted to be highly functional (by several noncoding scoring methods; Table 3-2, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t3-2>) and associated with some complex metabolic or hematological diseases (from GAD). Another associated SNP rs75595941, which is located in one intron of *DGRB3* (Adhesion G-protein-coupled receptor B3), has been associated with various complex diseases including narcolepsy (from GAD). Functional predictions of these top associated variants are listed in Table 3-2 (available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t3-2>).

### Gene-based analysis and pathway analysis

Gene-based analysis integrated weak SNP signals into a gene and thus increased signal-to-noise ratio in a genomic association study. We used a gene-based software program VEGAS2. The gene encoding an olfactory receptor, *OR11A1*, reached an uncorrected  $p = 2 \times 10^{-6}$  (but did not pass the genome-wide significance level of  $7.9 \times 10^{-7}$ ) at the discovery stage for BR. Among these candidate genes with uncorrected  $p < 0.05$  (Table 3-3, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t3-3>), *OR11L6* (encoding an olfactory receptor) reached corrected significance for NC [uncorrected  $p = 1.10 \times 10^{-5} < (0.05/3231) = 1.55 \times 10^{-5}$ ], whereas *MIR1178* was significant for VOLUNTARY (uncorrected  $p = 1.00 \times 10^{-6} < 1.55 \times 10^{-5}$ ) in the replication cohort.

We further performed pathway analysis to explore potential biological pathways involved in perceptual rivalry. No pathway was significant after multiple-testing correction at the discovery stage (significance threshold  $p = 1.71 \times 10^{-6}$ ), and none of the selected pathways (Table 3-4, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t3-4>) could be replicated (significance threshold  $p = 0.05/1936 = 2.58 \times 10^{-5}$ ). Two pathways, GO:0042992 (negative regulation of transcription factor import into nucleus) and GO:1901533 (negative regulation of hematopoietic progenitor cell differentiation), were shared by BR and NC with uncorrected  $p < 0.05$  at the discovery stage and also showed uncorrected  $p < 0.05$  for either BR or NC at the replication stage, suggesting their important role in spontaneous perceptual rivalry. It is of note that *PRMT1* (protein arginine methyltransferase 1) from pathway GO:1901533 was revealed with uncorrected  $p < 0.05$  at the gene level for both BR and NC at the discovery stage (Table 3-5, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t3-5>). The top associated gene *MIR1178* was involved in a candidate pathway, G13\_Signaling\_Pathway ( $p_{\text{discovery}} = 0.009$ ,  $p_{\text{replication}} = 0.001$ , without correction).

### Functional consequences revealed by brain anatomy

The two successfully replicated SNPs (association signals with FWER  $< 0.05$  at both the discovery and the replication stages) were tested for associations with brain anatomical features of interest by ANOVA, aiming for functional validation (Table 3-6, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t3-6>). Because 24 tests in total were conducted, the corrected significance threshold was  $p = 0.002$ . There was difference (uncorrected  $p = 0.04$ ) between the two rs184765639 genotypes (G/A and G/G) in the surface area of the left caudal middle frontal cortex (Table 3-6, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t3-6>). However, this difference was not significant after multiple-testing correction. This result suggests potential involvement of rs184765639 in the top-down modulation of perceptual rivalry. Furthermore, rs184765639 was significantly associated with voluntary modulation behavioral score with an uncorrected  $p = 0.041$  (note that only one SNP was tested for association with the behavioral score, VOLUNTARY; thus, this  $p$  value is significant (Table 3-6, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t3-6>)).

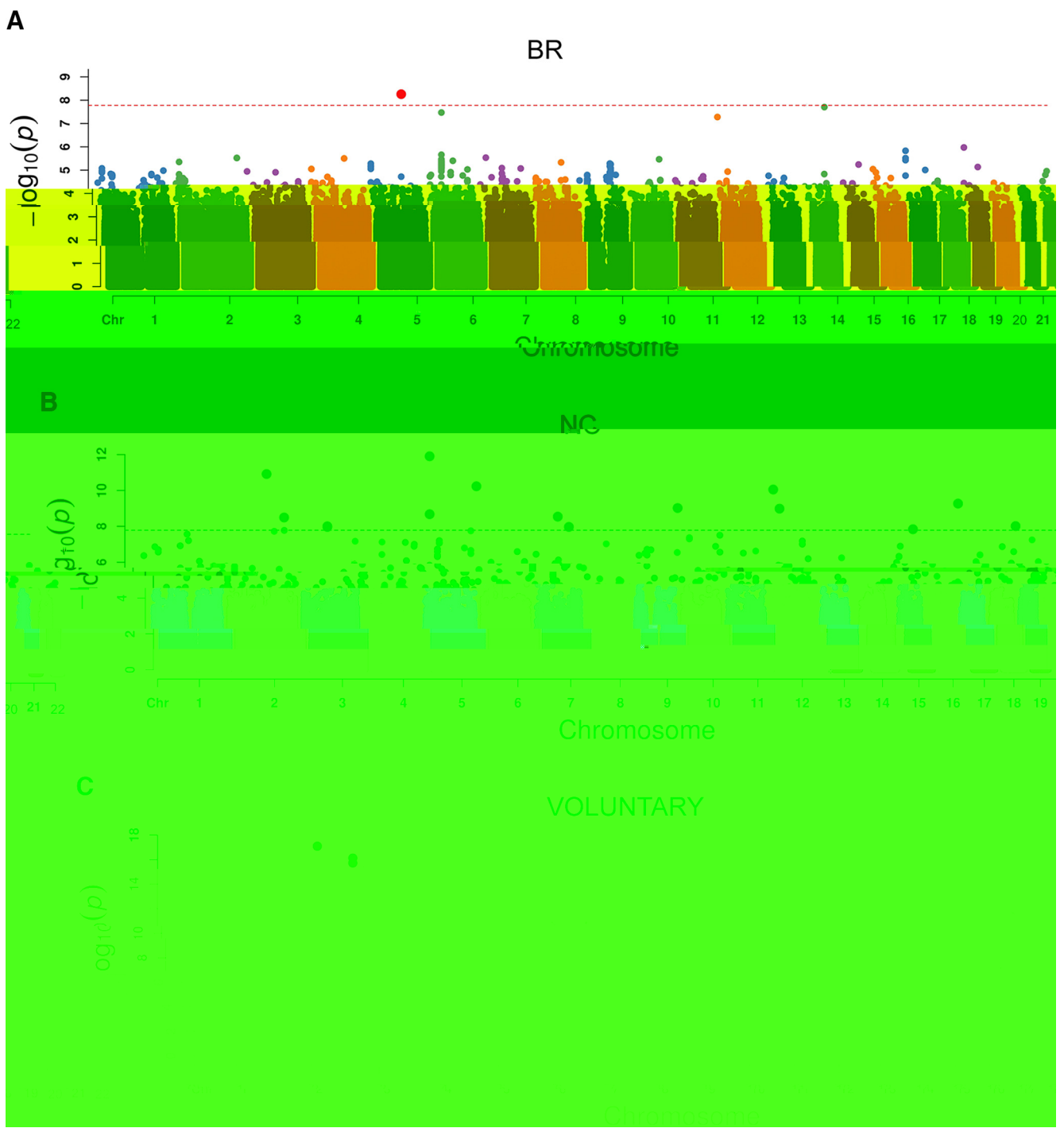
### Discussion

This study performed systematic genomic analyses at multiple levels to investigate an important phenomenon in human visual cognition, perceptual rivalry, and its voluntary modulation. Results from this study suggest several genes implicated in perceptual

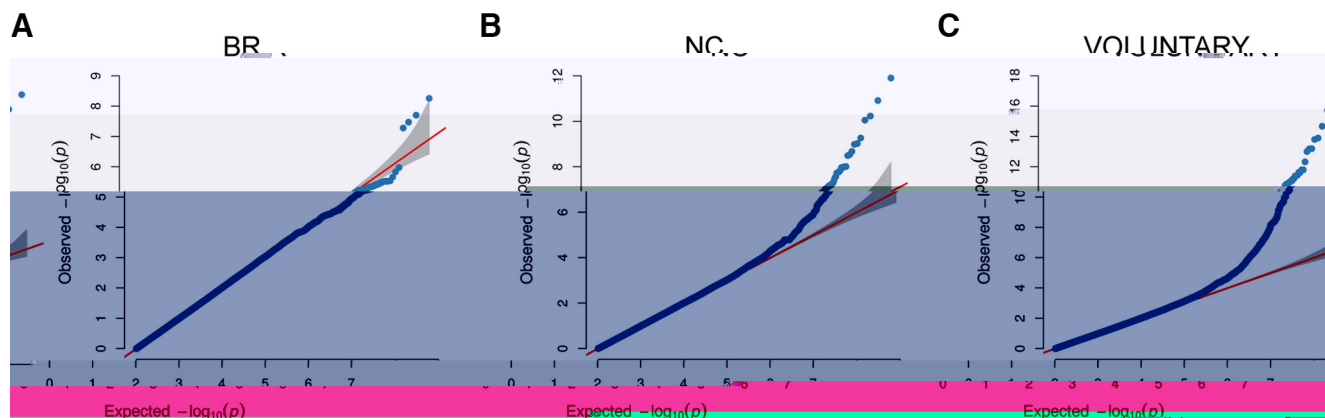
**Table 2. Estimations of heritability explained by all autosomal SNPs**

Phenotype	$\sigma_a^2$ (SE)	$\sigma_e^2$ (SE)	$\sigma_p^2$ (SE)	$h^2$ (SE)	$P$	$N$
BR	23.76 (12.67)	70.42 (12.60)	94.18 (2.47)	0.252 (0.134)*	0.026	2935
NC	2.06 (1.39)	6.84 (1.39)	8.90 (0.25)	0.231 (0.156)	0.069	2593
VOLUNTARY	0.00 (0.092)	0.588 (0.092)	0.588 (0.016)	0.000 (0.156)	0.5	2586

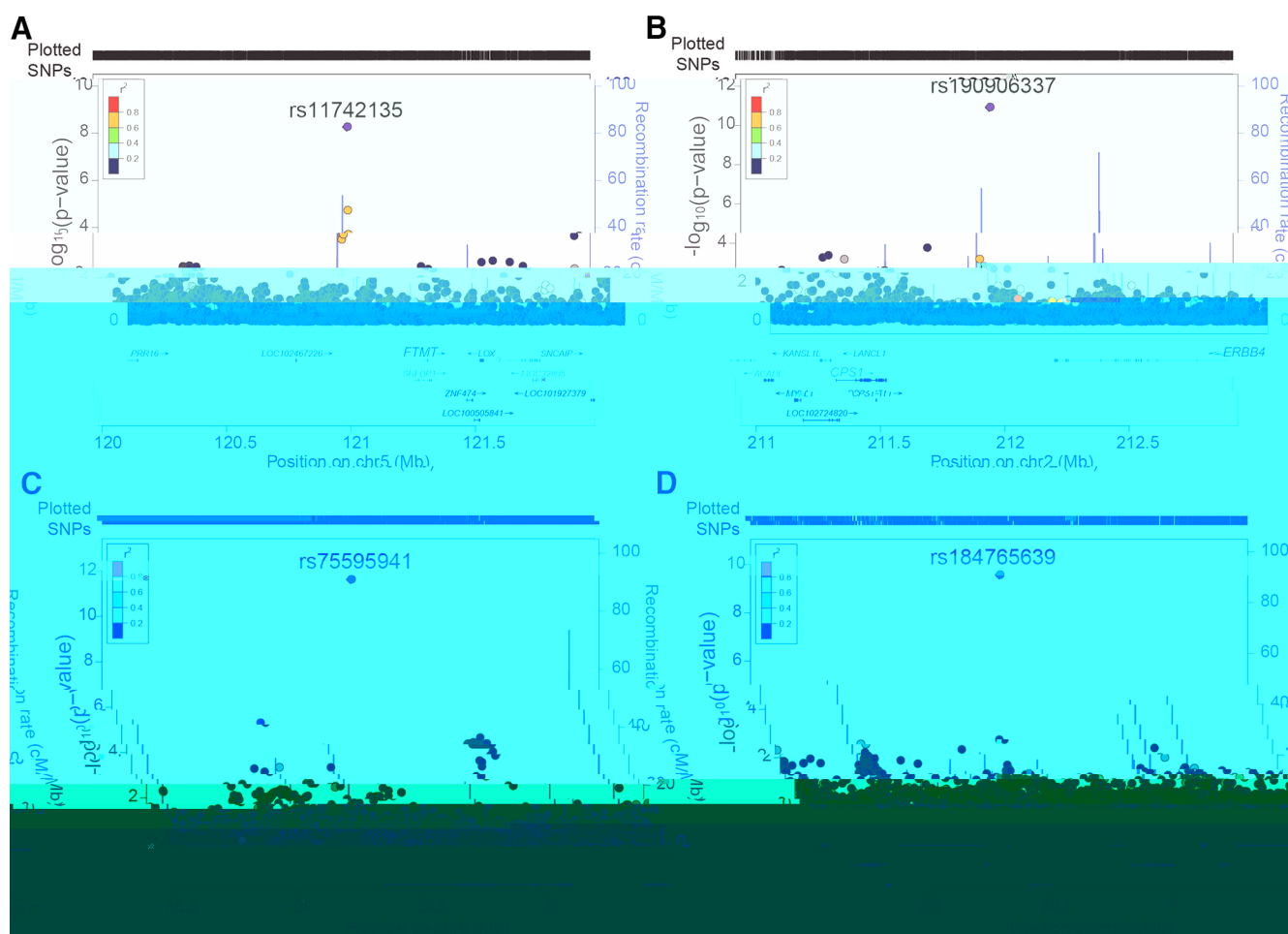
$\sigma_a^2, \sigma_e^2, \sigma_p^2$ , Estimates of variance explained by additive genetic effects, environmental effects, phenotype;  $h^2$ , heritability defined by  $\sigma_a^2/\sigma_p^2$ ;  $P$ ,  $p$  value for the estimation. BR, binocular rivalry; NC, Necker cube rivalry; VOLUNTARY, voluntary modulation strength. \* $p < 0.05$ . For genetic correlations, see Table 2-1 (available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018>).



**Figure 2.** Manhattan plots of genome-wide association study (GWAS) results at the discovery stage. GWAS results [ $-\log_{10}(p \text{ value})$ ] are shown in chromosomal order for individually genotyped quality-control-positive single nucleotide polymorphisms (SNPs) that were tested for linear regression with **(A)** perceptual switching rate in the binocular rivalry task (BR), **(B)** spontaneous perceptual switching rate with the Necker cube rivalry (NC), and **(C)** voluntary modulation strength (VOLUNTARY) in the discovery cohort. Results were based on an additive genetic model. Each dot represents one SNP. The red dotted line indicates the genome-wide significance level ( $1.67 \times 10^{-8}$ ). Genome-wide significant SNPs are shown as enlarged red dots. Chromosomes are shown in different colors for clarity.



**Figure 3.** Quantile-quantile plots of genome-wide association study (GWAS) results at the discovery stage. Observed  $p$  values relative to expected  $p$  values are plotted for (**A**) perceptual switching rate in the binocular rivalry task (BR), (**B**) spontaneous perceptual switching rate for the Necker cube rivalry (NC), and (**C**) voluntary modulation strength (VOLUNTARY), based on  $p$  values calculated using linear regression and with significant principal components included as covariates. Each blue dot represents one SNP. The red line indicates the null hypothesis of no association. The gray areas indicate 95% confidence intervals for the no association line.



**Figure 4.** Regional association plot for top associated loci. **A**, The genome-wide significant single nucleotide polymorphism (SNP) at the discovery stage for binocular rivalry (BR), rs11742135. **B**, The SNP with genome-wide significance at the discovery stage and  $p < 0.05$  at the replication stage for Necker cube rivalry (NC), rs190906337. **C**, One replicated SNP for voluntary modulation strength (VOLUNTARY), rs75595941. **D**, One replicated SNP for VOLUNTARY, rs184765639. Imputed genotypes from the discovery sample were considered for association tests. Genome build is hg19/1000 Genomes Nov 2014 ASN. Figures are plotted using the web-based LocusZoom program. Each dot represents one SNP.

rivalry and voluntary modulation of rivalry, and demonstrate that modern genomics, in addition to behavioral, electrophysiological, brain imaging, and brain stimulation methods, is an important new approach for studying human cognitive functions. We have performed genomic analyses at four levels: (1) SNP

heritability estimation from all SNPs across the genome, (2) SNP-based GWAS, (3) gene-based analysis, and (4) pathway analysis. In addition, we performed a genotypic differential analysis in the phenotype-related brain regions based on structural MRI data, which suggested a role of the SNPs identified from genomic anal-

**Table 3. GWAS results of replicated loci**

yses in regulating brain structural features involved in perceptual rivalry.

The GREML method estimates SNP heritability with assumptions different from those in twin studies. Our results showed that, ~25% of the phenotypic variance in spontaneous perceptual rivalry could be explained by common SNPs. The estimates for BR reached significance, validating the conclusion of non-zero heritability for perceptual rivalry. Previously, several twin studies explored genetic contributions to perceptual rivalry, and suggested the twin heritability of BR and NC to be ~50% (Miller et al., 2010; Shannon et al., 2011; Wang et al., 2014). It is approximately twice the SNP-based heritability measured in our study. The smaller heritability estimate via the SNP-based method than via twin-based studies could be due to the genotyped SNPs being only part of the overall genetic diversity. Contributions of many low-frequency variants were not measured. There may exist incomplete linkage between the genotyped markers and real causal variants (Yang et al., 2010). We also acknowledge that, because the test-retest reliability for the behavioral phenotypes was not systematically examined, potential random variations in our behavioral measurements might contribute to this lower SNP heritability.

It is possible that BR rate might take time to stabilize, though subjects had already taken 1 min training with the rivalry stimulus. Therefore, we performed additional analyses with the mean switching rate of only the last four experimental trials in the BR task. After quality control, two SNPs reached genome-wide significance level at the discovery stage, namely rs199954060 (chr12: 8985184delC,  $p = 3.4 \times 10^{-12}$ ) and rs144173227 (chr17: 79714232C>T,  $p = 2.7 \times 10^{-8}$ ), but they were not replicated. 73% of the SNPs with uncorrected  $p < 0.05$  for BR rate averaged across all the eight trials also showed uncorrected  $p < 0.05$  for that averaged across the last four trials. Furthermore, the relative ranking of  $p$  values from GWAS of the last four trials was similar to that from GWAS of all the eight trials (Spearman's correlation for the two sets of  $p$  values was 0.61). Also, there was a high correlation between the average switching rates for the first four trials and the last four trials ( $r = 0.81$ ). However, this correlation is not so high as that in Miller et al. (2010) between their blocks 2 and 3 ( $r = 0.93$ ), suggesting that some degree of switching rate stabilization might occur in the first few trials of BR viewing.

The successfully replicated SNP rs75595941 for VOLUNTARY is located in an intron of *ADGRB3*, a p53-target gene that encodes a brain-specific angiogenesis inhibitor and regulates synaptogenesis and dendritic spine formation (Kaur et al., 2003; Lanoue et al., 2013). *ADGRB3* is overexpressed in the frontal cortex (from the Human Integrated Protein Expression Database), which is involved in top-down control (Johnston et al., 2007). *ADGRB3* has been linked to psychiatric disorders such as schizophrenia (Lips et al., 2012) and bipolar disorder (McCarthy et al., 2012). Several genes suggested by multiple levels of genomic

analyses, such as *PRMT1*, are also involved in neuronal development.

Recently, there has been a genomic study on BR in which researchers performed a single GWAS with <1000 participants and indicated several SNPs with  $p$  values between  $1 \times 10^{-7}$  and  $1 \times 10^{-5}$  (but without a replication study; Bosten et al., 2015). Here, our study investigated BR and NC as well as voluntary modulation of NC in a larger sample, and we have found genome-wide significant variants that survived replication. We conducted a meta-analysis of our study and Bosten et al.'s (2015) study with all the 12 SNPs they suggested to be associated with BR ( $p < 1 \times 10^{-5}$ ). No significant associations for the 12 SNPs were found in our study although three SNPs reached uncorrected  $p = 0.06$ . The failure to replicate all the SNPs found by Bosten et al. (2015) might be due to the shorter test duration in their task and different populations in their study and ours.

It has been proposed that BR and NC may share some common biological mechanisms (Andrews and Purves, 1997; Logothetis, 1998; Carter and Pettigrew, 2003). Based on our data, we believe that the common mechanisms are quite limited. First, these two kinds of perceptual rivalry were only weakly, albeit significantly, correlated phenotypically ( $r = 0.19$ ). This is consistent with some previous studies (Gallagher and Arnold, 2014), but differs from Shannon et al. (2011) who reported a correlation coefficient as high as 0.47. Second, there were only a few shared genes (Table 3-5, available at <https://doi.org/10.1523/JNEUROSCI.1970-17.2018.t3-5>) and pathways revealed by the genomic analyses.

A major limitation of this study is the small sample size (by the standards of genetic research), which reduces the power to detect association signals with a relatively small effect size and may lead to inaccurate SNP heritability estimation. So the suggestive findings from this study need replication with much larger sample sizes. A second limitation is the gender imbalance in the sample, which might introduce a sex factor into genetic associations. It would be better to recruit a more gender-balanced cohort for future replication studies. A third limitation is that mixed perceptual states commonly occur in the BR task but they were not recorded in the current study. Before the analyses, we excluded those participants who experienced persistent mixed percepts. The remaining participants were instructed to classify mixed percept into either perceptual state based on which image occupied the majority of the viewed stimulus field. This response protocol differed from the twin study by Miller et al. (2010) and this difference may have affected our heritability and association results. Future research can examine the relevance of recording and excluding mixed perceptual states and whether the duration of mixed percepts might itself be a parameter of interest in clinical conditions (as appears to be the case in autism; see Introduction). A fourth limitation is that subjects were only instructed to increase their perceptual switching rate for measuring voluntary



modulation. In this regard, it is noteworthy that Hunt and Guilford (1933) reported the difference between Necker cube switching rates in bipolar versus control subjects was even greater when subjects were instructed to slow down the perceptual reversal (i.e., the opposite of our instruction). We chose to focus on the speed up instruction to avoid possible confounding factors associated with the slow down instruction, but future studies examining the genetics underlying voluntary modulation of perceptual rivalry could assess the slow down process. A fifth limitation is that VOLUNTARY phenotype shows no SNP heritability but has genome-wide significant threshold SNPs. The reason for this conundrum could be due to the relatively small sample size and the low imputation INFO score for genome-wide significant SNPs that might smear associations and result in no consolidated peaks (Fig. 4). Future investigation on the heritability of VOLUNTARY is needed.

Perceptual rivalry has been implicated as a potential endophenotype for bipolar disorder (Pettigrew and Miller, 1998; Ngo et al., 2011; Vierck et al., 2013). Several genes associated with perceptual rivalry in our study are suggested to be associated with bipolar disorder by literature, such as BR-related gene *FE2L2* (Rizak et al., 2014) and NC-related gene *SRI* (Beasley et al., 2006), which further supports the relationship between perceptual rivalry and bipolar disorder. Investigation of the genetics of perceptual rivalry may contribute to our understanding of molecular mechanisms underlying psychiatric disorders. Deficits in voluntary control were found in many psychological disorders such as bipolar disorder (Blumberg et al., 2003; Dickstein et al., 2004). Several candidate genes associated with voluntary modulation strength in our study have also been implicated in these disorders, such as *DGRB3* (McCarthy et al., 2012). This raises the possibility that measuring voluntary modulation strength through

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